

A stochastic convergence analysis for Tikhonov-Regularization with sparsity constraints

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- Introduction
- A convergence theorem
- Convergence rates
- Numerical examples



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Can we prove convergence (-rates) for Tikhonov-Regularization with sparsity-penalty if instead of $\|y - y^\delta\| \leq \delta$ an explicit stochastic error model is used?

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with $\mathbf{A} \in \mathcal{L}(\mathcal{X}, \mathcal{Y})$ where \mathcal{X} and \mathcal{Y} are Hilbert spaces.

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$$\|\mathbf{Ax} - \mathbf{y}^\delta\|^2 + \hat{\alpha} \Phi_{\mathbf{w},p}(\mathbf{x}) \rightarrow \min_{\mathbf{x}} \quad (1)$$

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- Penalty $\Phi_{\mathbf{w},p}(\mathbf{x}) = \sum_{\lambda \in \Lambda} w_\lambda |\langle \mathbf{x}, \psi_\lambda \rangle|^p$ for an ONB $\{\psi_\lambda\}$

- computation requires discretization, done via projections

$$P_m : \mathcal{Y} \rightarrow \mathbb{R}^m, \quad \mathbf{y} \mapsto y, \quad \text{e.g. point evaluation}$$

$$T_n : \mathcal{X} \rightarrow \mathbb{R}^n, \quad x = T_n \mathbf{x} = \{\langle \mathbf{x}, \psi_i \rangle\}_{i=1, \dots, n}$$

where $\{\psi_i\}_{i=1}^{\infty}$ is ONB in \mathcal{X} .

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- each component of y carries *stochastic* noise, $y^\sigma = y + \varepsilon$, $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_m)$.
- Define $A := P_m \mathbf{A} T_n^*$, then we want to find x s.t.

$$Ax = y^\sigma \tag{2}$$

We use Bayes' formula

to calculate the solution. In this framework, every quantity is treated as a random variable in a complete probability space $(\Omega, \mathcal{F}, \mathbb{P})$.

$$\pi_{post}(x|y^\sigma) = \frac{\pi_\varepsilon(y^\sigma|x)\pi_{pr}(x)}{\pi_{y^\sigma}(y^\sigma)}.$$

- $\pi_{post}(x|y^\sigma)$ posterior density
- $\pi_\varepsilon(y^\sigma|x)$ likelihood function
- $\pi_{pr}(x)$ prior distribution
- $\pi_{y^\sigma}(y^\sigma)$ data distribution (irrelevant)

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gaussian error model:

$$\pi_\varepsilon \propto \exp\left(-\frac{1}{2\sigma^2} \|Ax - y^\sigma\|^2\right),$$



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Reasons:

- "easy" characterization with coefficients of a wavelet expansion
- sparsity-promoting properties
- discretization invariance (Lassas, Saksman, Siltanen '09)

Definition of the wavelet expansion

Let $\phi \in C^{\tilde{s}}(\mathbb{R})$ be a scaling function and $\psi \in C^{\tilde{s}}(\mathbb{R})$ a compactly supported wavelet, satisfying

$$\begin{aligned}\phi_{j,k}(t) &= 2^{\frac{j}{2}} \phi(2^j t - k), \\ \psi_{j,k}(t) &= 2^{\frac{j}{2}} \psi(2^j t - k), \quad j, k \in \mathbb{Z}.\end{aligned}$$

such that the collection $\{\psi_{j,k} : j, k \in \mathbb{Z}\}$ constitutes an orthonormal basis of $L_2(\mathbb{R})$ and for each $h \in L_2(\mathbb{R})$ we have

$$h = \sum_{k=-\infty}^{\infty} \langle f, \phi_{0,k} \rangle \phi_{0,k} + \sum_{j=0}^{\infty} \sum_{k=-\infty}^{\infty} \langle f, \psi_{j,k} \rangle \psi_{j,k}.$$

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assumption:

only finitely many nonzero scalar products on each scale

- with this, a d -dimensional basis can be constructed



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- let $\{\psi_\lambda : \lambda \in \Lambda\}$ denote the set of all wavelets ψ , also including the scaling functions ϕ where Λ is an appropriate index set, possibly infinite
- set $|\lambda| = j$, then
- $\mathbf{x} \in B_p^s(\mathbb{R}^d) \subset L^2(\mathbb{R}^d)$, $s < \tilde{s}$, if

$$\|\mathbf{x}\|_{s,p} := \left(\sum_{\lambda \in \Lambda} \underbrace{2^{\varsigma p |\lambda|}}_{w_\lambda} |\langle \mathbf{x}, \psi_\lambda \rangle|^p \right)^{1/p} < \infty$$

and $\varsigma = s + d(\frac{1}{2} - \frac{1}{p}) \geq 0$. We focus on $1 \leq p \leq 2$.

Besov-space random variables

Definition (adapted from Lassas/Saksman/Siltanen, 2009)

Let $1 \leq p < \infty$ and $s \in \mathbb{R}$. Let $(X_\lambda^\alpha)_{\lambda \in \Lambda}$ be independent identically distributed real-valued random variables with probability density function

$$\pi_X(\tau) = c_p^\alpha \exp\left(-\frac{\alpha|\tau|^p}{2}\right), \quad c_p^\alpha = \left(\frac{\alpha}{2}\right)^{\frac{1}{p}} \frac{p}{2\Gamma(\frac{1}{p})}, \quad \tau \in \mathbb{R}.$$

Let U be the random function

$$U(t) = \sum_{\lambda \in \Lambda} 2^{-s|\lambda|} X_\lambda^\alpha \psi_\lambda(t), \quad t \in \mathbb{R}^d.$$

Then we say U is distributed according to a B_p^s -prior, $U \propto \exp\left(-\frac{\alpha}{2} \|U\|_{s,p}^p\right)$.



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Theorem (adapted from Lassas/Saksman/Siltanen, 2009)

Let U be as before, $2 < \alpha < \infty$ and take $r \in \mathbb{R}$. Then the following three conditions are equivalent:

- (i) $\|U\|_{B_p^r(\mathbb{R}^d)} < \infty$ almost surely,
- (ii) $\mathbb{E} \exp\left(\|U\|_{B_p^r(\mathbb{R}^d)}^p\right) < \infty$,
- (iii) $r < s - \frac{d}{p}$.

same result as [LSS 2009], but here \mathbb{R}^d instead of \mathbb{T}^d considered

How to avoid this phenomenon?

- “finite model” (MI)

- consider discretization level m and n fixed, finite index set Λ_n
- Then

$$X_n(t) := \sum_{\lambda \in \Lambda_n} 2^{-c|\lambda|} X_\lambda^\alpha \psi_\lambda(t) \Rightarrow \|X_n\|_{s,p}^p = \sum_{\lambda \in \Lambda_n} |X_\lambda^\alpha|^p < \infty$$

- and $\mathbb{P}(\|X_n\|_{s,p} > \varrho) = \frac{\Gamma(\frac{n}{p}, \frac{\alpha \varrho^p}{2})}{\Gamma(\frac{n}{p})} \leq \frac{1}{\varrho} \sqrt[p]{\frac{2n}{\alpha p}}$

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- define $X(t)$ in $B_p^r(\mathbb{R}^d)$ with $s < r - \frac{d}{p}$, then
 - $\mathbb{E}(\|X\|_{B_p^s(\mathbb{R}^d)}) = \left(\frac{2}{\alpha p} \left(c_\lambda^1 + c_\lambda^2 \sum_{j=0}^{\infty} 2^{-j((r-s)p-d)} \right) \right)^{\frac{1}{p}} < \infty$
 - and $\mathbb{P}(\|X\|_{s,p} > \varrho) \leq \frac{1}{\varrho} \mathbb{E}(\|X\|_{B_p^s(\mathbb{R}^d)})$

Recall

$$\pi_{post}(x|y^\sigma) = \frac{\pi_{pr}(x)\pi_\varepsilon(y^\sigma|x)}{\pi_{y^\sigma}(y^\sigma)}.$$

$\pi_\varepsilon(y^\sigma|x)$ gaussian noise, $\pi_{pr}(x)$ Besov-Space prior

$$\Rightarrow \pi_{post}(x|y^\sigma) \propto \exp\left(-\frac{1}{2\sigma^2}\|Ax - y^\sigma\|^2\right) \cdot \exp\left(-\frac{\alpha}{2}\|T_n^*x\|_{s,p}^p\right)$$

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we are interested in the *maximum a-priori* solution

$$x_{\text{map}} = \operatorname{argmax}_{x \in \mathbb{R}^n} \pi_{post}(x|y^\sigma)$$

or equivalently

$$x_\alpha^{\text{map}} = \operatorname{argmin}_{x \in \mathbb{R}^n} \left\|Ax - y^\sigma\right\|^2 + \alpha\sigma^2\|T_n^*x\|_{B_p^s(\mathbb{R}^d)}^p \quad (3)$$

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same functional as in deterministic case, but $\|y - y^\sigma\| \leq \delta$ does not hold



- stochastic setting requires different measure for convergence
- convergence in expectation may be too strict
- instead, we use the *Ky Fan metric*

Definition

Let x_1 and x_2 be random variables in a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ with values in a metric space $(\mathcal{X}, d_{\mathcal{X}})$. The distance between x_1 and x_2 in the *Ky Fan metric* is defined as

$$\rho_K(x_1, x_2) := \inf\{\epsilon > 0 : \mathbb{P}(d_{\mathcal{X}}(x_1(\omega), x_2(\omega)) > \epsilon) < \epsilon\}.$$

- allows combination of deterministic and stochastic quantities

Ky Fan error estimate

Theorem (Neubauer, Pikkarainen, 2008)

Let ξ be a random variable with values in \mathbb{R}^m . Assume that the distribution of ξ is $\mathcal{N}(y_0, \sigma^2 I)$ with $\sigma > 0$. Then it holds in $(\mathbb{R}^m, \|\cdot\|)$ that

$$\rho_K(\xi, y_0) \leq \min \left\{ 1, \sqrt{2} \sigma \sqrt{m - \ln^- \left(\sigma^2 2\pi m^2 \left(\frac{e}{2} \right)^m \right)} \right\},$$

where $f^-(h) := \min\{0, f(h)\}$.

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where $f^-(h) := \min\{0, f(h)\}$.

in practice \ln -term mostly inactive, then

$$\rho_K(\xi, y_0) \leq \min \left\{ 1, \sqrt{2}\sigma \sqrt{m} \right\},$$

c.f. $\mathbb{E}(\|\xi - y_0\|) = \sigma \sqrt{m}$

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Let x^\dagger be the unique solution of the equation $Ax = y$ with minimum value of $\Phi(\cdot)$.

Theorem (adapted from Hofinger, '06)

Let $\alpha, \sigma > 0$, $1 \leq p \leq 2$, $\|A\| < 1$ and $N(A) = 0$ for $p = 1$. Assume $w_\lambda \geq c > 0 \forall \lambda \in \Lambda$. Let x_α^{map} be the solution of (3). If $\alpha = \alpha(\sigma)$ is chosen such that $\hat{\alpha} = \alpha\sigma^2 \rightarrow 0$ and $\frac{|\ln \sigma|}{\alpha} \rightarrow 0$ as $\sigma \rightarrow 0$, then

$$\lim_{\sigma \rightarrow 0} \rho_K(x_\alpha^{map}, x^\dagger) = 0.$$

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as long as $\sigma^2 2\pi m^2 \left(\frac{\epsilon}{2}\right)^m > 1$, then $\alpha \rightarrow \infty$ is sufficient

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as long as $\sigma^2 2\pi m^2 \left(\frac{\epsilon}{2}\right)^m > 1$, then $\alpha \rightarrow \infty$ is sufficient
 main idea for the proof: use Ky Fan metric and split
 $\Omega = \Omega_{\det}(\sigma) \cup \Omega_{\text{unbound}}(\sigma)$



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deterministic convergence rate, DDD '04

Assume \mathbf{A} fulfils, for all $h \in L^2$

$$A_l^2 \sum_{\lambda} 2^{-2|\lambda|\beta} |\langle h, \psi_{\lambda} \rangle|^2 \leq \|Ah\|^2 \leq A_u^2 \sum_{\lambda} 2^{-2|\lambda|\beta} |\langle h, \psi_{\lambda} \rangle|^2. \quad (4)$$

and $\|\mathbf{x}^{\dagger}\|_{s,p} \leq \varrho$, $\varrho > 0$. Then

$$\begin{aligned} \sup\{\|\mathbf{x}_{\alpha}^{\text{map}} - \mathbf{x}\| : \mathbf{x} \in \mathcal{X}, \mathbf{y} \in \mathcal{Y}, \|\mathbf{A}\mathbf{x} - \mathbf{y}\| \leq \delta, \|\mathbf{x}\|_{s,p} \leq \varrho\} \\ < C \left(\frac{\delta + \delta'}{A_l} \right)^{\frac{\varsigma}{\beta + \varsigma}} (\varrho + \varrho')^{\frac{\beta}{\beta + \varsigma}} \end{aligned}$$

with $\delta' = (\delta^2 + \hat{\alpha}\varrho^p)^{\frac{1}{2}}$ and $\varrho' = (\varrho^p + \frac{\delta^2}{\hat{\alpha}})^{\frac{1}{p}}$.

Theorem

Let A fulfil (4) and assume that we have an a-priori estimate $\|x^\dagger\|_{s,p} \leq \varrho$ for some $\varrho > 0$. Set $a_m := \ln\left(\frac{2^m}{2\pi m^2}\right)$. Then as $\sigma \rightarrow 0$, x_α^{map} converges with the parameter choice $\alpha = \alpha(\sigma, \varrho, \beta, \varsigma, p, m, n)$ fulfilling

$$f(\alpha) := \min \left\{ 1, 2 \left(\frac{\sqrt{2}}{A_l} \sigma \sqrt{a_m - 2 \ln \sigma + \frac{\alpha \varrho^p}{2}} \right)^{\frac{\varsigma}{\beta + \varsigma}} \left(\left(\varrho^p + \frac{2}{\alpha} (a_m - 2 \ln \sigma) \right)^{1/p} \right)^{\frac{\beta}{\beta + \varsigma}} \right\} \\ - \frac{\Gamma\left(\frac{m}{2}, m\right)}{\Gamma\left(\frac{m}{2}\right)} - \mathbb{P}(\|x\|_{B_p^s} > \varrho) = 0$$

to the unique solution x^\dagger and

$$\rho_K(x_\alpha^{map}, x^\dagger) = \mathcal{O} \left(\left(\sigma \sqrt{1 + |\ln \sigma| + \alpha \varrho^p} \right)^{\frac{\varsigma}{\beta + \varsigma}} \left(\left(\varrho^p + \frac{1 + |\ln \sigma|}{\alpha} \right)^{1/p} \right)^{\frac{\beta}{\beta + \varsigma}} \right).$$

where $\mathbb{P}(\|x_n\|_{B_p^s} > \varrho) = \frac{\Gamma\left(\frac{n}{p}, \frac{\alpha \varrho^p}{2}\right)}{\Gamma\left(\frac{n}{p}\right)}$ or $\mathbb{P}(\|x\|_{B_p^s} > \varrho) = \frac{\mathbb{E}\|x\|_{B_p^s}}{\varrho}$



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- We consider a convolution problem

$$[Ax](s) = [k * x](s) = \int_{\mathbb{R}^d} k(s-t)x(t)dt, \quad s \in \mathbb{R}^d \quad (5)$$

- using a kernel

$$\widehat{k}(\xi) = \frac{c_{\kappa,\beta}}{(1 + \kappa|\xi|^2)^{\beta/2}}, \quad \xi \in \mathbb{R}^d, \quad c_{\kappa,\beta} \text{ s.t. } \|\widehat{k}\|_{L_2(\mathbb{R}^d)} < 1$$

- thus $\|A\| < 1$ and (4) is fulfilled with chosen β
- $p = 1, d = 1$

Iteration [Daubechies, De Mol, Defrise 2004]:

With $x_0 = 0$,

$$x_{k+1} = \mathcal{S}_{\mathbf{w},p}(x_k + A^*(y^\sigma - Ax_k)), \quad k = 1, 2, \dots,$$

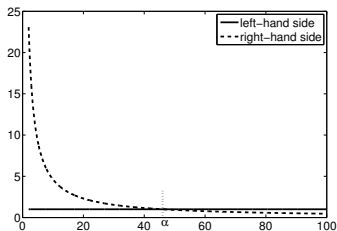
where $\mathcal{S}_{\mathbf{w},p}(h) := \sum_{\lambda \in \Lambda} S_{w_\lambda,p}(\langle h, \psi_\lambda \rangle) \psi_\lambda$ is defined component-wise ($p = 1$) via

$$S_{\omega,1}(\xi) := \begin{cases} \xi - \frac{\varepsilon}{2} & \text{if } \xi \geq \frac{\varepsilon}{2} \\ 0 & \text{if } |\xi| < \frac{\varepsilon}{2} \\ \xi + \frac{\varepsilon}{2} & \text{if } \xi \leq -\frac{\varepsilon}{2} \end{cases} .$$

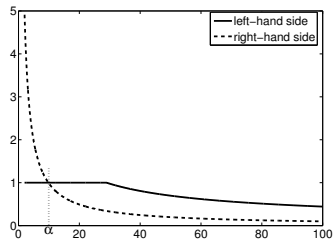
Parameter choice rule illustrated

$$\sigma = 0.01, m = 2500, \varsigma = 0.5, \beta = 1, \varrho = 2.16$$

model (MI), $s = 1$



model (MII), $s = 1, r = 2$



example of a solution

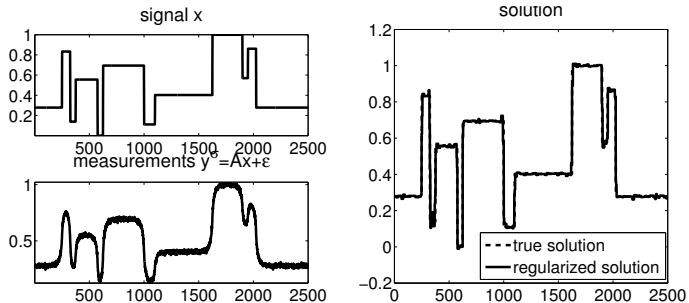


Figure : (MI), $\sigma = 0.01$, exact ϱ , $s = 1$, $\beta = 1$. $\alpha = 45.85$
 $\Rightarrow \hat{\alpha} = \alpha\sigma^2 = 0.004585$

comparison of (MI) and (MII), m, n fixed, $\sigma \rightarrow 0$

all plots averaged over 20 individual simulations

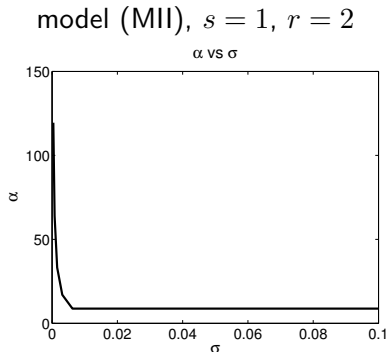
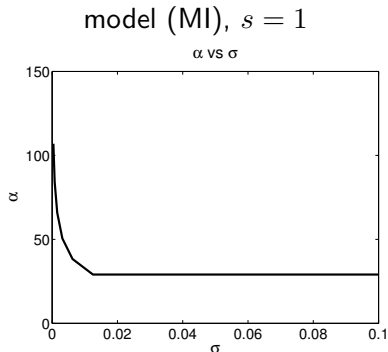


Figure : α plotted against σ , $n = m = 2500$, $\beta = 1$, exact ρ

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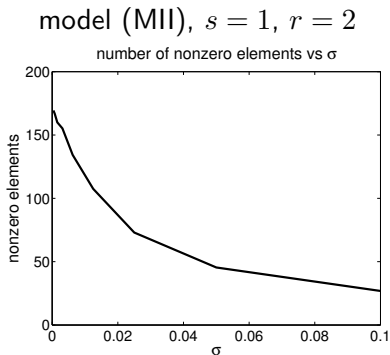
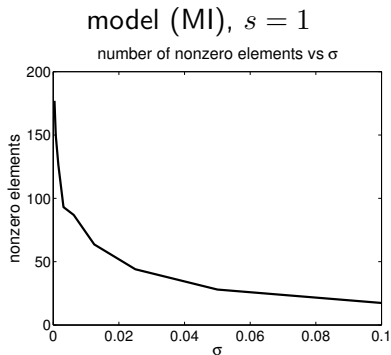


Figure : number of recovered nonzero coefficients plotted against σ ,
 $n = m = 2500, \beta = 1$, exact ϱ

comparison of (MI) and (MII), m, n fixed, $\sigma \rightarrow 0$

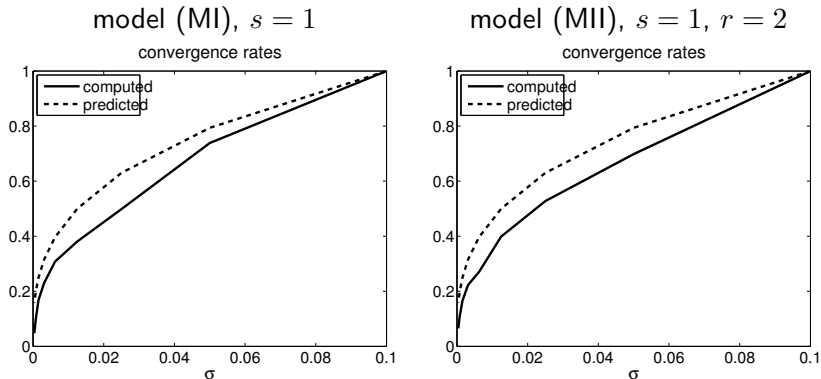
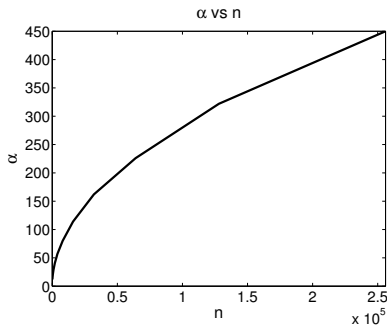


Figure : predicted and observed convergence rates plotted against σ ,
 $n = m = 2500$, $\beta = 1$, exact ϱ

comparison of (MI) and (MII), σ fixed, m, n variable

model (MI), $s = 1$



model (MII), $s = 1, r = 2$

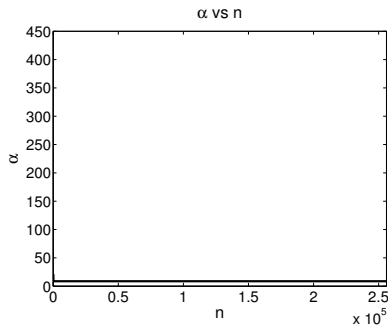


Figure : α plotted against n , $\sigma = 0.01$, $\beta = 1$, exact ϱ

comparison of (MI) and (MII), σ fixed, m, n variable

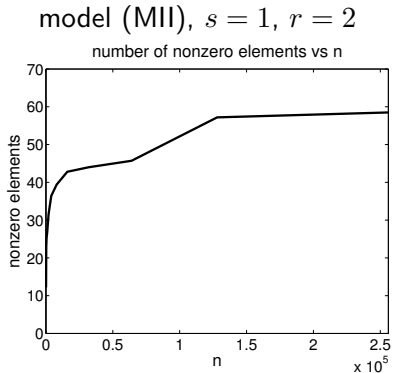
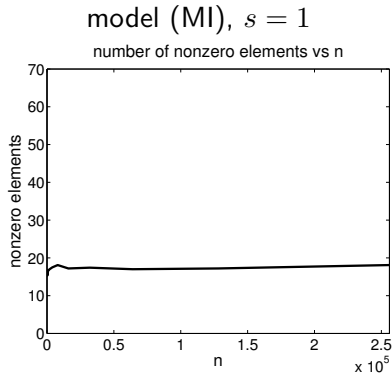
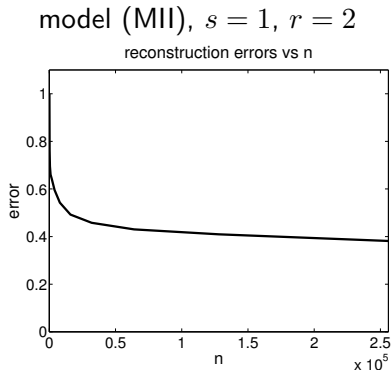
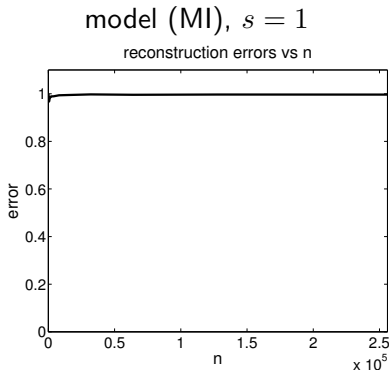


Figure : number of recovered nonzeros plotted against n , $\sigma = 0.01$, $\beta = 1$, exact ρ

comparison of (MI) and (MII), σ fixed, m, n variableFigure : reconstruction error plotted against n , $\sigma = 0.01$, $\beta = 1$, exact ϱ

A 2D convolution example

$$\sigma = 0.1, \beta = 1, \alpha = 130.5, \hat{\alpha} = 1.3$$

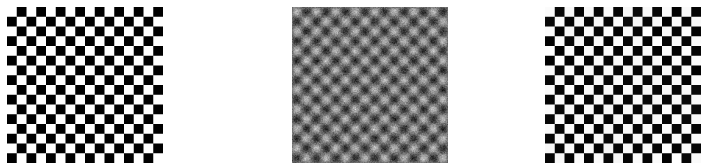







Figure : true solution - measurements - recovered solution

exactly the 68 original coefficients were reconstructed



-  G., R. Ramlau, *A stochastic convergence analysis for Tikhonov-Regularization with sparsity constraints*, in preparation
-  M. Lassas, E. Saksman, S. Siltanen, *Discretization-invariant Bayesian inversion and Besov space priors*, Inverse Probl. Imaging, 2009.
-  I. Daubechies, M. Defrise, C. De Mol, *An iterative thresholding algorithm for linear inverse problems with a sparsity constraint*, Comm. Pure Appl. Math. 57, 2004.
-  A. Neubauer, H.K. Pikkarainen, *Convergence results for the Bayesian inversion theory*, J. Inverse Ill-Posed Probl. 16, 2008.
-  J. Kaipio, E. Somersalo, *Statistical and computational inverse problems*, Springer-Verlag, New York, 2005.

Thank you for attention! Are there questions?